

Is employment growth affected by the introduction of broadband services? Evidence from firms in Ireland

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Abstract

This paper examines the effects of deploying digital subscriber line (DSL) broadband on job creation within existing firms. We use spatial information on broadband and firm locations, exploiting geographical and temporal variation in broadband availability in Ireland during 2007-2014. This is linked to a panel of firm-level data on employment and other characteristics. Econometric models are used to explore the relationship between DSL and employment over time and across local economic and industry contexts. We also investigate whether effects might vary depending upon local educational attainment. We find little evidence of a general effect of DSL roll-out on employment in Irish firms.

Keywords: broadband; employment growth; Ireland

JEL Classification Codes: D22, L96, R12

1. Introduction

The provision of high-speed internet services necessitates significant capital investment, which is often publicly funded. It is thus important that research strives to provide clarity on the potential economic benefits of broadband diffusion. In a recent survey of the literature, Bertschek et al. (2015) highlight that the expansion of broadband has been associated with various economic benefits including GDP growth, increased firm presence, and productivity gains. There is, however, notable heterogeneity in these findings, with many results sensitive to context.

The economic gains from the introduction of broadband may also manifest as changes in

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local employment growth. This is reflected in a selection of empirical studies including Czernich et al. (2011) and Ivus and Boland (2015). The former analyse the effect of digital subscriber line (DSL) availability on the German municipality level unemployment rate. A positive but quantitatively small association is identified in OLS regressions but the effect is not robust to modelling with instrumental variables and thus cannot be considered causal. The latter find that DSL roll-out promoted both aggregate employment growth and average wages across rural Canada. This effect was, however, limited to services industries and was most pronounced among sub-sectors which could be classified as “IT-intensive”. In urban areas, broadband roll-out showed signs of curtailing growth, suggesting that the technology was of greatest benefit to service firms who needed to overcome geographical constraints in their business models.

If broadband availability does affect employment, the effect may be heterogeneous across firms and local economic landscapes. Education is likely a mediating factor. Broadband access may be complementary to highly skilled labour but act as a substitute to positions requiring lower skill levels (Akerman et al. 2015). In order to test this “skill complementarity” hypothesis, some studies present empirical models which interact broadband availability with the proportion of the local population who have obtained a degree or higher level of education. Atasoy (2013) and Kolko (2012) are two examples. The former identifies a positive effect of DSL on U.S. county-level employment that varies significantly in magnitude across the distribution of the education proxy. The latter finds that the proportion of degree holders over the age of 25 mediates a positive effect of DSL on employment in the early years of rollout. The effect dissipates in later samples possibly indicating a complementarity between broadband and highly skilled labour and specialised knowledge only when the technology was new.

The current research contributes to what remains a relatively sparse literature focusing on the relationship between firm-level employment and DSL broadband. DSL is of particular research interest as it was the predominant technology used by SMEs in Ireland who had access to the internet during the early years of broadband roll-out. We use detailed information on the roll-out of DSL in the Republic of Ireland to exploit geographical and temporal variation in broadband availability between the years 2007-2014. By combining this information with a panel of geocoded firm-level employment data, we are able to study broadband diffusion micro-econometrically with the firm as the unit of analysis. Such dis-aggregation is uncommon to date. We also contribute to the existing literature by further investigating the potential variation of the economic effects of broadband across firm size, sector and type. Furthermore, we empirically test the skill complementarity hypothesis in the case of Ireland by studying the interactions between broadband diffusion and a spatial measure of human capital.

2. Methods

For research purposes, DSL is conventionally characterised in one of two ways: availability or adoption. We focus on availability. This choice is made for two reasons. First, despite a potential role for demand response to the new technology, it is the presence of the broadband infrastructure that is the essential enabling factor in any link between DSL and employment. Second, extending the availability of high-quality broadband access to underserved regions has, to date, been the primary focus of most public policies that aim to improve DSL penetration. We allow for potential time delays between the date on which DSL services are initially enabled and that by which firms have effectively embedded the new technology into their business processes by using a lagged representation of DSL availability. Specifically, DSL enters our empirical models as the share of the previous year for which DSL services were available to a given firm.

To investigate whether local educational attainment moderates the effect of DSL on employment, we use two strategies. First, following the convention of Kolko (2012) and Atasoy (2013) we interact DSL availability with the proportion of local residents who have obtained a univer-

sity degree or higher level of education. Second, we split our sample at the median local educational attainment among firms that became DSL enabled during our sample to perform subgroup analyses.

Our dependent variable measures annual proportional change in the number of employees at each firm. By representing the absolute change as a proportion of total employment in the previous year, we account for the likelihood that the magnitude of employment changes vary systematically by firm size. In order to fully encompass all changes into a single measure, we use weighted full-time equivalent (FTE) employment. The variable weights every two part-time workers in our data as a single full-time employee.

We thus estimate a set of fixed effects panel regressions of the following form:

$$\Delta Y_{it} = \alpha + \beta_1 DSL_{it-1} + \beta_2 (DSL_{it-1} \times DEG) + \beta_3 X_{it} + \eta_i + \epsilon_{it} \quad (1)$$

ΔY_{it} denotes proportional change in FTE employment for firm i in time t with respect to time $t - 1$. DSL_{it-1} is our variable measuring broadband availability. DEG is the proportion of the local Electoral Division (ED) who have obtained a university degree or higher level of education. X_{it} is a vector of other relevant controls including year effects that account for unobserved time-specific factors, η_i is a firm fixed effect which captures time-invariant heterogeneity among firms and ϵ_{it} is an error term such that $\epsilon_{it} \sim IID(0, \sigma_\epsilon^2)$. The interaction term is, naturally, omitted in education subgroup analyses.

While there are some concerns in the literature that the relationship between DSL and employment could suffer from endogeneity because of reverse causal or omitted factor mechanisms (Czernich 2014; Ivus and Boland 2015; Kolko 2012), they are not shared here. Central to this are three points: First, our dependent variable is an annualised flow in employment. We argue that, even under a reverse causality scenario, the commercial (or policy) decision to deploy DSL to a particular region is likely to be influenced more by the local population density than expected future employment growth rates. Second, as noted, we characterise DSL diffusion by its availability rather than its adoption. It is more plausible that omitted firm-level factors drive changes in broadband adoption than it is that they affect the national pattern of DSL deployment. Third, we lag DSL availability by a year in our regressions, so the model relates current employment growth to past DSL levels. Nevertheless, we suggest that the results reported in this paper should be interpreted cautiously and do not support strong claims about causality.

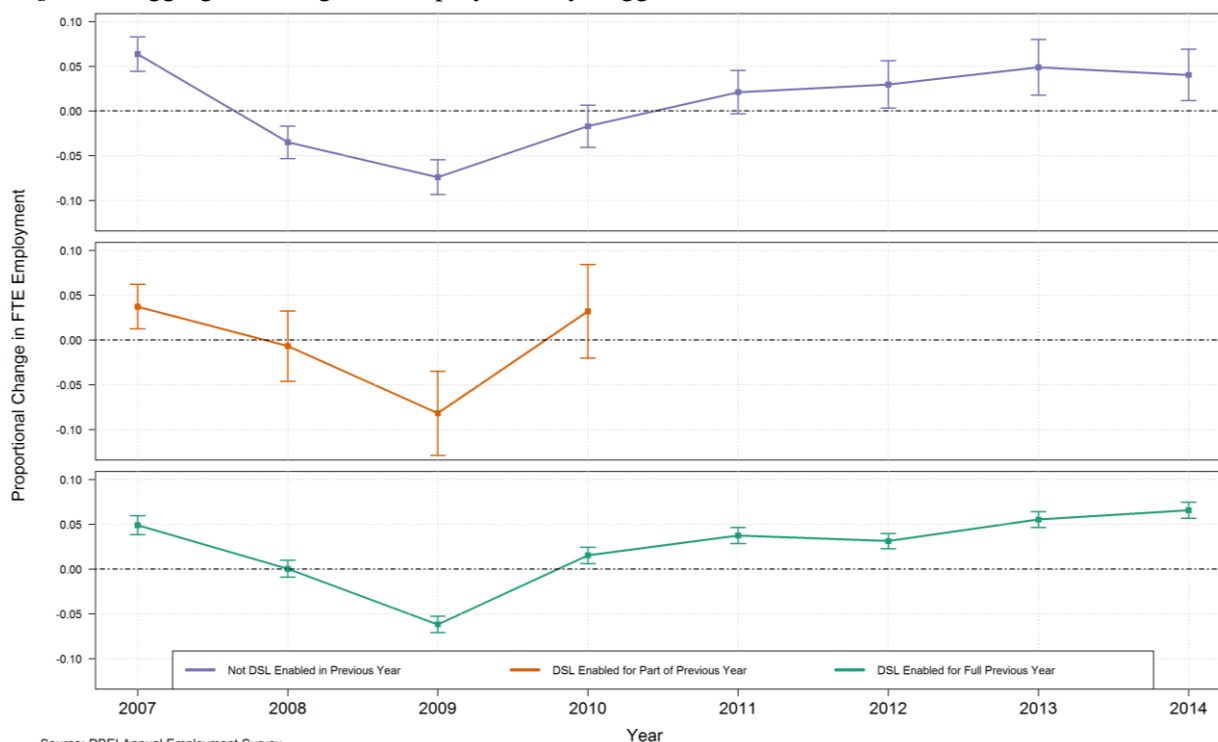
3. Data

The data used combines information from three sources. The first is the Annual Employment Survey (AES), carried out by the Department of Business, Enterprise and Innovation (DBEI) within the Irish Government. The dataset is a firm-level panel created from an annual census of all manufacturing and service firms under the remit of the Irish State's enterprise development agencies. We consider employment changes for existing firms in the years 2007-2014. As such, firm entries to or exits from the dataset are excluded.¹ Observations for which the dependent variable is greater than 3 standard deviations above its mean are considered to be outliers and are also excluded. The effective sample thus comprises of 40,693 firm-year observations. For each observation, the AES provides the geocoded location of the firm, a NACE Rev.2 code which outlines the sector in which the firm operates, whether the firm is under Irish or foreign ownership as well as the number of full and part-time workers employed. Figure 1 plots employment changes aggregated by our DSL availability variable over time. While job losses are

¹ Where zero FTE employees are reported for a single year during a period of sustained employment, this is assumed to be an error. FTE employment is imputed as the mean of its value from the preceding and following years.

observed in years that coincide with the global recession, the overall trend does not appear to differ significantly by DSL status. More detailed descriptive statistics appear in Table 1.

Figure 1. Aggregate Changes in Employment by Lagged DSL Status.



Source: DBEI Annual Employment Survey.
 Note: The series in the centre panel is incomplete in order to preserve firm anonymity. Error bars represent 95% confidence intervals.

Table 1. Descriptive statistics.

	Proportional Change in FTE Employment								
	Full Sample			Below Median Education			Above Median Education		
	N	Mean	SD	N	Mean	SD	N	Mean	SD
NACE Rev. 2 Sub-Sector									
Agriculture, Forestry & Fishing (A)	844	0.014	0.294	416	0.017	0.300	428	0.010	0.288
Mining and Quarrying (B)	426	-0.012	0.255	221	-0.002	0.250	205	-0.023	0.260
Manufacturing (C)	24,695	0.009	0.280	10,090	0.004	0.281	14,605	0.011	0.278
Water, Sewage & Waste Mgmt. (E)	346	0.041	0.358	91	-0.016	0.339	255	0.062	0.363
Construction (F)	887	-0.005	0.329	348	-0.010	0.320	539	-0.002	0.335
Wholesale & Retail Trade (G)	796	0.001	0.269	314	-0.011	0.239	482	0.008	0.286
Transportation and Storage (H)	488	0.016	0.288	346	0.029	0.292	142	-0.015	0.276
Accommodation & Food Services (I)	277	-0.002	0.379	157	0.004	0.395	120	-0.010	0.357
Information and Communication (J)	5,983	0.067	0.366	1,390	0.047	0.351	4,593	0.073	0.371
Financial & Insurance Activities (K)	426	0.081	0.258	196	0.050	0.235	230	0.107	0.274
Professional Scientific & Technical (M)	1,492	0.056	0.361	430	0.046	0.362	1,062	0.060	0.362
Admin. and Support Services (N)	902	0.048	0.343	351	0.041	0.349	551	0.052	0.338
Education (P)	429	0.035	0.313	119	0.052	0.366	310	0.029	0.291
Health & Social Work (Q)	340	0.044	0.339	134	0.021	0.392	206	0.060	0.299
Arts, Entertainment & Rec. (R)	1,174	0.019	0.374	425	-0.002	0.353	749	0.031	0.385
Other Services (S)	933	0.000	0.330	562	0.001	0.332	371	-0.001	0.328
Other*	255	0.031	0.423	109	0.051	0.455	146	0.015	0.399
Firm Ownership Status									
Irish Owned	35,531	0.020	0.313	14,190	0.012	0.307	21,341	0.025	0.316
Foreign Owned	5,162	0.029	0.271	1,509	0.006	0.244	3,653	0.038	0.281

DSL Status in Previous Year									
No DSL	5,760	0.007	0.316	2,767	0.004	0.312	2,993	0.009	0.320
DSL Enabled (Part Year)	1,098	0.011	0.302	548	0.000	0.294	550	0.021	0.309
DSL Enabled (Full Year)	33,835	0.024	0.306	12,384	0.013	0.299	21,451	0.030	0.310
Total	40,693			15,699			24,994		

Source: Department of Business Enterprise and Innovation - Annual Employment Survey.

Note: *Remaining NACE sectors have been aggregated in order to preserve anonymity.

We use spatial information to assess whether broadband services were available to each firm in each year. We combine the firm geocodes with panel data that maps the availability of DSL through the network of fixed-line local telephone exchanges across Ireland. This dataset, which was provided to the authors by the country’s primary network operator Eir, includes the locations of all 1,125 local telephone exchanges in the country, the boundaries of the area which each one serves as well as the date on which each was enabled for DSL service. The nature of DSL technology means that service quality is a function of line distance from the exchange and service is not available if the line distance exceeds 5,000m. In many cases, this means that service remains unavailable in areas of enabled exchanges. We thus define the boundary of enabling as a 5,000m buffer zone around a given exchange or the designated exchange area boundary, whichever is closer to the exchange location. Geographic Information Software (GIS) allows us to map out the effective enabled region under this definition. The results of this GIS process are mapped in Figure 2 while Figure 3 highlights how DSL expansion relates to firms in the AES over time.

Figure 2. Development of DSL enabled regions.

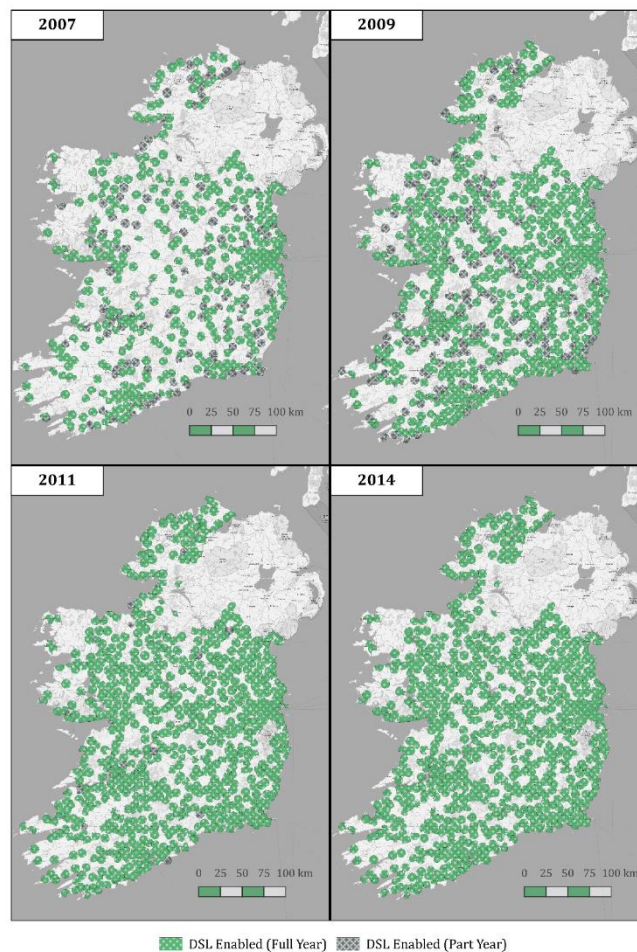
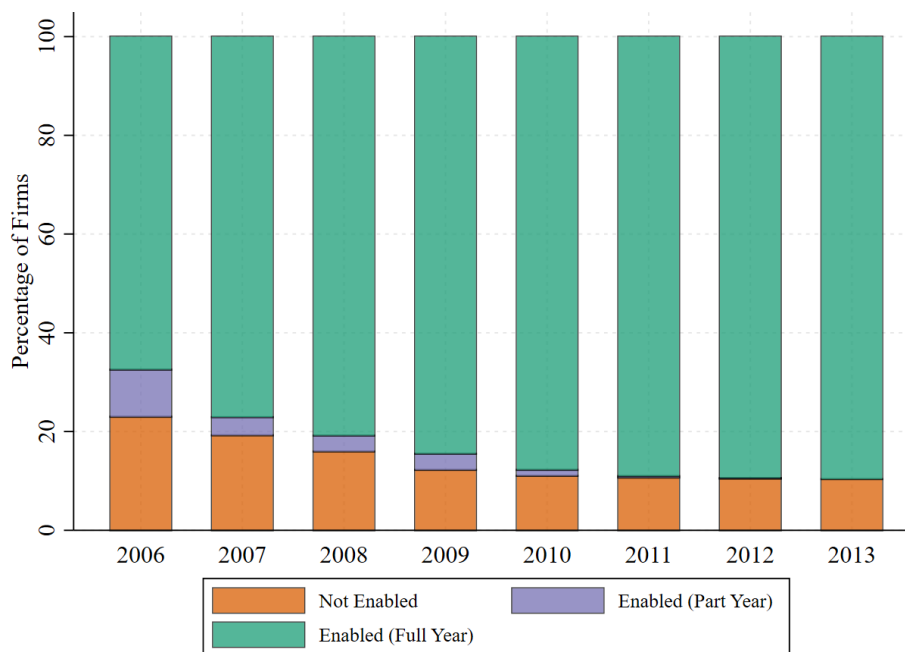


Figure 3. Firms enabled for DSL (2006-2013).



Our final data source, the Census small area population data from 2011, is provided by the Central Statistics Office of Ireland (CSO). From this dataset, we extract the proportion of the population of a given area with a third level degree. The level of aggregation used is the local Electoral Division (ED) of which there are 3,409 in the Republic of Ireland. The geographic area of an ED is the most likely of all available data options to be representative of a local labour market. GIS software is again used to map each firm to its local ED.

Since complete DSL enabling occurred in Ireland's capital, Dublin long before our period of analysis, the county is excluded from the study. Furthermore, Dublin has a uniquely skilled labour market and a density of economic activity great enough to engender spatial agglomerations and knowledge spill-overs not present elsewhere. These fundamental differences further motivate the exclusion.

4. Results

Table 2 presents results from our main fixed effects models. We do not find statistically significant effects of DSL availability on employment growth rates. While the sign on the DSL coefficient is positive (as expected) in the full sample specification, statistical significance is only observed at the 10% level. The coefficient on the interaction between DSL and educational attainment is also not statistically significant. Indeed, when the sample is split by education subgroup, the DSL coefficient is not statistically significant in either sub-sample. The values of the DSL coefficients remain relatively unchanged when foreign-owned firms are excluded but no statistical significance is observed. Consistent with Figure 1, one broadly observes negative year effects, relative to 2012, in years coinciding with the global recession, albeit that statistical significance is sporadic. These specifications present little evidence that changes in employment were linked to DSL availability in our sample.

To investigate heterogeneity in the results across firm types, we use Eurostat/NACE technological intensity ratings. The ratings are a standard, pan-European tool used to group firms whose production processes require particular levels of knowledge and technology. A breakdown of the six relevant categories along with a sample list of associated economic activities is outlined in Appendix A. The potential for DSL effects to be mediated by education is assessed

only through interaction terms in these models as there are insufficient observations for sub-group analyses. The results are presented in Table 3.

Table 2. Fixed effects panel regressions.

	Dependent Variable: Proportional Change in FTE Employment					
	Full Sample			Irish Owned Firms		
	(1) All	(2) LED ^a	(3) HED ^b	(4) All	(5) LED ^c	(6) HED ^d
<i>Lagged DSL Status</i>	0.0444* (0.0262)	0.0214 (0.0139)	-0.00415 (0.0143)	0.0408 (0.0280)	0.0208 (0.0145)	-0.0102 (0.0155)
<i>DSL Lag × Degree</i>	-0.279 (0.198)			-0.279 (0.215)		
Year Effects						
2007	0.0592*** (0.00614)	0.0539*** (0.00978)	0.0628*** (0.00789)	0.0598*** (0.00673)	0.0543*** (0.0106)	0.0638*** (0.00874)
2008	-0.00360 (0.00597)	-0.00438 (0.00967)	-0.00329 (0.00759)	-0.00658 (0.00654)	-0.00844 (0.0104)	-0.00551 (0.00843)
2009	-0.0705*** (0.00579)	-0.0656*** (0.00953)	-0.0737*** (0.00728)	-0.0695*** (0.00635)	-0.0645*** (0.0103)	-0.0730*** (0.00807)
2010	-0.00388 (0.00583)	-0.0122 (0.00931)	0.00141 (0.00748)	-0.00598 (0.00639)	-0.0158 (0.0100)	0.000580 (0.00829)
2011	0.0123** (0.00588)	0.00196 (0.00949)	0.0188** (0.00749)	0.0114* (0.00641)	-0.00106 (0.0101)	0.0196** (0.00828)
2012	[ref]	[ref]	[ref]	[ref]	[ref]	[ref]
2013	0.0143** (0.00592)	0.0148 (0.00978)	0.0140* (0.00744)	0.0220*** (0.00652)	0.0179* (0.0106)	0.0245*** (0.00827)
2014	0.0101* (0.00573)	0.0195** (0.00972)	0.00443 (0.00707)	0.0162** (0.00632)	0.0238** (0.0105)	0.0113 (0.00787)
<i>Constant</i>	0.0186* (0.0111)	-0.00617 (0.0137)	0.0284** (0.0138)	0.0189 (0.0117)	-0.00454 (0.0143)	0.0296** (0.0149)
<i>Fixed Effects</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Observations</i>	40,693	15,699	24,994	35,531	14,190	21,341
<i>Firms</i>	6,517	2,468	4,049	5,723	2,244	3,479
<i>Within R2</i>	0.0151	0.0138	0.0164	0.0149	0.0137	0.0163

Notes: ^aLower education sample (below median for firms with DSL status change). ^bHigher education sample (above median for firms with DSL status change). Errors allow for clustering at the firm level. *p<0.1; **p<0.05; ***p<0.01

In services sectors, we find no statistically significant relationship between DSL, or its interaction with education, and employment. Across manufacturing sectors, the only result which returns with statistical significance is that of the high-technology manufacturing (HTM) category. Here, DSL shows signs of promoting employment growth. Offsetting this, however, is a larger negative coefficient on the education interaction which suggests that rather than complementing high skilled labour, broadband is acting as a substitute. While surprising in the context of the existing literature, it may be intuitive that among this group of high technology firms, labour could be substituted with technological solutions. A significant caveat to this result is that relatively few firms in this sub-sample experience a change in DSL availability in our period of analysis. This may suggest the observed significance is an artefact of our data.

Appendix B includes other specifications aimed at testing the stability of the coefficients. First, we test for heterogeneity across firm sizes with a split in the sample at a median of 20 FTE employees. Similarly, we present a specification that includes only data pertaining to the years prior to 2012. After this time, the variation in DSL availability was significantly reduced since the deployment of first generation broadband had been completed. We also address the possibility that firms in urban and rural areas may respond differently to the expansion of DSL

availability. We divide the sample into two by geography. An ‘urban’ category includes firms in all counties that contain an urban centre designated as a city, and all other firms are categorised as ‘non-urban’. The DSL coefficients lack statistical significance under each of these checks.

Table 3. Fixed effects models by Eurostat-NACE technological intensity ratings.

	Dependent Variable: Proportional Change in FTE Employment					
	Manufacturing				Services	
	(1)	(2)	(3)	(4)	(5)	(6)
	LTM ^a	MLTM ^b	MHTM ^c	HTM ^d	LKIS ^e	KIS ^f
<i>Lagged DSL</i>	-0.0248 (0.0425)	-0.00520 (0.0603)	0.00508 (0.0739)	0.623*** (0.229)	-0.00882 (0.105)	0.0555 (0.0618)
<i>DSL Lag × Degree</i>	0.265 (0.331)	0.168 (0.512)	-0.0189 (0.501)	-3.474*** (1.315)	0.297 (0.608)	-0.537 (0.399)
Year Effects						
2007	0.0394*** (0.00986)	0.0577*** (0.0139)	0.0120 (0.0157)	0.0492* (0.0282)	0.0515** (0.0239)	0.105*** (0.0154)
2008	-0.0349*** (0.00989)	-0.0166 (0.0135)	-0.0469*** (0.0160)	0.0151 (0.0273)	0.0283 (0.0247)	0.0352** (0.0138)
2009	-0.0671*** (0.00985)	-0.122*** (0.0132)	-0.129*** (0.0159)	-0.0572** (0.0234)	-0.0293 (0.0230)	-0.0439*** (0.0131)
2010	-0.0231** (0.00985)	-0.00916 (0.0139)	-0.0220 (0.0172)	0.0452 (0.0274)	-0.0211 (0.0204)	0.0242* (0.0131)
2011	-0.00923 (0.00976)	0.0196 (0.0141)	0.00987 (0.0171)	-0.00922 (0.0261)	-0.0182 (0.0209)	0.0400*** (0.0131)
2012	[ref]	[ref]	[ref]	[ref]	[ref]	[ref]
2013	0.00584 (0.00988)	0.0311** (0.0141)	0.0180 (0.0187)	-0.0181 (0.0235)	0.00127 (0.0241)	0.00994 (0.0125)
2014	0.0122 (0.00950)	0.0521*** (0.0136)	-0.0141 (0.0157)	-0.0511** (0.0217)	0.0166 (0.0210)	-0.00479 (0.0132)
<i>Constant</i>	0.00629 (0.0160)	-0.0224 (0.0253)	0.0382 (0.0290)	0.0372 (0.0662)	-0.0163 (0.0323)	0.0741** (0.0293)
<i>Fixed Effects</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Observations</i>	11,463	6,480	4,647	2,105	3,202	10,200
<i>Firms</i>	1,776	989	712	333	504	1,789
<i>Within R2</i>	0.0168	0.0454	0.0328	0.0212	0.00694	0.0148

Notes: ^aLow technology manufacturing; ^bMedium-low technology manufacturing; ^cMedium-high technology manufacturing; ^dHigh technology manufacturing; ^eLess knowledge intensive services; ^fKnowledge intensive services. Errors clustered at firm level. *p<0.1; **p<0.05; ***p<0.01

5. Concluding remarks

We find little evidence of any association between DSL rollout and employment growth within existing Irish firms during this period, even when educational attainment is taken into account. This finding is in contrast with evidence from other jurisdictions suggesting that broadband tends to be complementary with highly educated labour, further emphasising the context-sensitivity of the effects broadband has on firms. Our analysis has some limitations that must be noted. The reliance on Euclidean distance to assess the availability of DSL likely implies at least some measurement error. The accuracy of our proxy in mirroring the exact locations of underground cable infrastructure cannot be tested in the data. Similarly, our measure of educational attainment is based on area averages. This may not fully reflect the skill-sets of a particular firm’s employees as workers could be chosen from a labour pool that is not bounded geographically by the local electoral division or characterised by its spatial averages. These measures do, however, represent the best feasible approximations given the currently available

data. Furthermore, we cannot be certain as to whether or not any firms in our sample had internet access through technologies other than DSL. However, as evidenced by additional survey data presented in Figure C1 (Appendix C), the majority of enterprises who used a broadband service in Ireland at the time did so through a DSL connection.

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Appendix A – Eurostat – NACE technology ratings

Table A1. Eurostat – NACE technology ratings.

Technological Intensity	Code	Sample Economic Activities
<i>Knowledge-intensive services</i>	KIS	Water transport; Air transport; Publishing activities; Programming and broadcasting activities; Computer programming, consultancy and related activities; Telecommunications; Financial and insurance activities; Security and investigation activities; Public administration; Education; Human Health.
<i>Less knowledge-intensive intensive services</i>	LKIS	Wholesale and retail trade; Accommodation and food service activities; Repair of motor vehicles and motorcycles; Real estate, rental and leasing activities; Office administration and office support; Travel agency, tour operator reservation services.
<i>High-technology manufacturing</i>	HTM	Manufacture of pharmaceutical products; Manufacture of computer, electronic and optical products.
<i>Medium-high-technology manufacturing</i>	MHTM	Manufacture of chemicals and chemical products; Manufacture of electrical equipment; Manufacture of motor vehicles.
<i>Medium-low-technology manufacturing</i>	MLTM	Manufacture of coke and refined petroleum products; Manufacture of rubber and plastic; Repair and installation of machinery and equipment.
<i>Low-technology manufacturing</i>	LTM	Manufacture of food products and beverages; Manufacture of textiles and apparel; Reproduction of recorded media; Manufacture of paper products; Manufacture of furniture; Other manufacturing.

Appendix B – Additional specifications

Table B1. Fixed effect panel regressions by firms size.

	Dependent Variable: Proportional Change in FTE Employment					
	Median FTE < 20			Median FTE > 20		
	(1)	(2)	(3)	(4)	(5)	(6)
	All	LED ^a	HED ^b	All	LED ^a	HED ^b
<i>Lagged DSL Status</i>	0.0370 (0.0327)	0.0249 (0.0171)	-0.00228 (0.0187)	0.0662 (0.0409)	0.0175 (0.0234)	-0.00828 (0.0186)
<i>DSL Lag x Degree</i>	-0.206 (0.252)			-0.462 (0.290)		
Year Effects						
2007	0.0669*** (0.00813)	0.0641*** (0.0128)	0.0689*** (0.0105)	0.0432*** (0.00850)	0.0326** (0.0135)	0.0495*** (0.0109)
2008	-0.00152 (0.00784)	-0.00371 (0.0124)	-0.000211 (0.0101)	-0.00741 (0.00846)	-0.00353 (0.0146)	-0.0102 (0.0103)
2009	-0.0574*** (0.00757)	-0.0542*** (0.0122)	-0.0597*** (0.00964)	-0.0986*** (0.00819)	-0.0906*** (0.0141)	-0.104*** (0.0100)
2010	-0.00520 (0.00759)	-0.0156 (0.0120)	0.00150 (0.00980)	-0.000572 (0.00836)	-0.00394 (0.0136)	0.00139 (0.0106)
2011	0.0153** (0.00764)	-0.00562 (0.0121)	0.0288*** (0.00987)	0.00556 (0.00844)	0.0200 (0.0145)	-0.00291 (0.0104)
2012	[ref]	[ref]	[ref]	[ref]	[ref]	[ref]
2013	0.0221*** (0.00778)	0.0200 (0.0128)	0.0234** (0.00981)	-0.00223 (0.00826)	0.00317 (0.0137)	-0.00555 (0.0104)
2014	0.0200*** (0.00775)	0.0248* (0.0130)	0.0171* (0.00962)	-0.0104 (0.00715)	0.00837 (0.0125)	-0.0215** (0.00867)
<i>Constant</i>	0.00661 (0.0140)	-0.0137 (0.0166)	0.0160 (0.0177)	0.0434*** (0.0163)	0.00727 (0.0240)	0.0551*** (0.0190)
<i>Fixed Effects</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Observations</i>	27,978	10,969	17,009	12,715	4,730	7,985
<i>Firms</i>	4,711	1,802	2,909	1,806	666	1,140
<i>Within R2</i>	0.0121	0.0121	0.0129	0.0301	0.0268	0.0333

Notes: ^aLower education sample (below median for firms with DSL status change). ^bHigher education sample (above median for firms with DSL status change). Errors allow for clustering at the firm level. *p<0.1; **p<0.05; ***p<0.01.

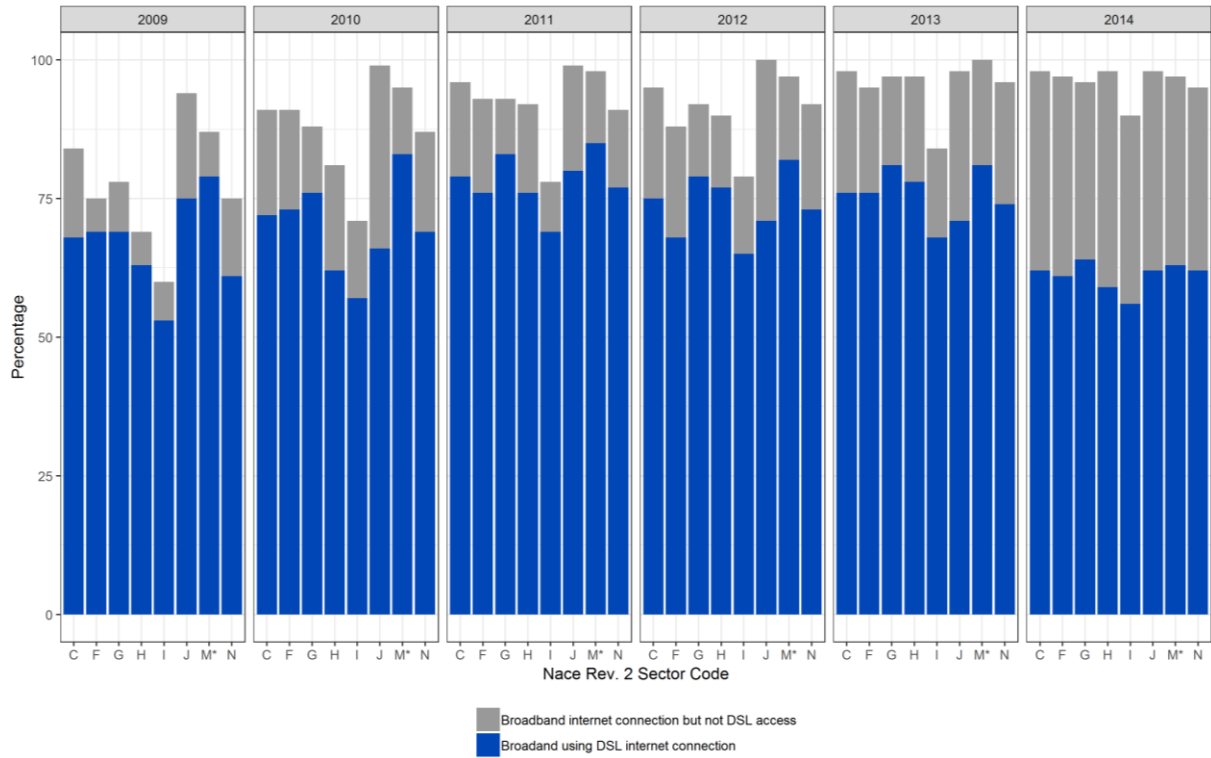
Table B2. DSL Coefficients from Additional Models.

	Lagged DSL Status	SE	Lagged DSL x Degree	SE	Obs
Pre-2012 Observations					
Full Sample	0.0344	(0.0322)	-0.206	(0.234)	25,922
Below Median Education Level	0.0159	(0.0183)			10,130
Above Median Education Level	0.000858	(0.0196)			15,792
Urban Vs. Non-urban Counties					
Urban Counties					
All	0.0699	(0.0496)	-0.242	(0.310)	14,809
Below Median Education Level	0.0472	(0.0344)			3,646
Above Median Education Level	0.0260	(0.0233)			11,163
Non-urban Counties					
All	0.0516	(0.0326)	-0.478*	(0.273)	25,884
Below Median Education Level	0.0119	(0.0141)			12,053
Above Median Education Level	-0.0271	(0.0179)			13,831

Notes: Errors are clustered at the firm level. *p<0.1; **p<0.05; ***p<0.01

Appendix C – Broadband connections in Irish firms

Figure C1. Enterprises in Ireland who report having a DSL connection.



Source: CSO Information Society Statistics - Enterprises

Note: * Category M (Professional, Scientific & Technical) here excludes veterinary activities.